





# 1 ExpFamilyPCA.jl: A Julia Package for Exponential 2 Family Principal Component Analysis

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## 6 Summary

7 Principal component analysis (PCA) ([Hotelling, 1933](#); [Jolliffe, 2002](#); [Pearson, 1901](#)) is popular  
8 for compressing, denoising, and interpreting high-dimensional data, but it underperforms on  
9 binary, count, and compositional data because the objective assumes data is normally distributed.  
10 Exponential family PCA (EPCA) ([Collins et al., 2001](#)) generalizes PCA to accommodate data  
11 from any exponential family distribution, making it more suitable for fields where these data  
12 types are common, such as geochemistry, marketing, genomics, political science, and machine  
13 learning ([Greenacre, 2021](#); [Hastie et al., 2009](#)).

14 ExpFamilyPCA.jl is a library for EPCA written in Julia, a dynamic language for scientific  
15 computing ([Bezanson et al., 2012](#)). It is the first EPCA package in Julia and the first in any  
16 language to support EPCA for multiple distributions.

## Statement of Need

18 EPCA is used in reinforcement learning ([Roy et al., 2005](#)), sample debiasing ([R. Huang &  
19 Lee, 2023](#)), and compositional analysis ([Gan & Valdez, 2024](#)). Wider adoption, however,  
20 remains limited due to the lack of implementations. The only other EPCA package is written  
21 in MATLAB and supports just one distribution ([Chambrier, 2016](#)). This is surprising, as other  
22 Bregman-based optimization techniques have been successful in areas like mass spectrometry  
23 ([Nozaki & Nakamoto, 2017](#)), ultrasound denoising ([J. Huang & Yang, 2013](#)), topological  
24 data analysis ([Edelsbrunner & Wagner, 2019](#)), and robust clustering ([Banerjee et al., 2005](#)).  
25 These successes suggest that EPCA holds untapped potential in signal processing and machine  
26 learning.

27 The absence of a general EPCA library likely stems from the limited interoperability between  
28 fast symbolic differentiation and optimization libraries in popular languages like Python and  
29 C. Julia, by contrast, uses multiple dispatch which promotes high levels of generic code  
30 reuse ([Karpinski, 2019](#)). Multiple dispatch allows ExpFamilyPCA.jl to integrate fast symbolic  
31 differentiation ([Gowda et al., 2022](#)), optimization ([Mogensen & Riseth, 2018](#)), and numerically  
32 stable computation ([Mächler, 2015](#)) without requiring costly API conversions. As a result,  
33 ExpFamilyPCA.jl delivers speed, stability, and flexibility, with built-in support for most common  
34 distributions (§ [Supported Distributions](#)) and flexible constructors for custom distributions (§  
35 [Custom Distributions](#)).

36 **Principal Component Analysis**

37 **Geometric Interpretation**

38 Given a data matrix  $X \in \mathbb{R}^{n \times d}$  with  $n$  observations and  $d$  features, PCA seeks the closest  
 39 low-rank approximation  $\Theta \in \mathbb{R}^{n \times d}$  by minimizing the reconstruction error

$$\begin{aligned} & \underset{\Theta}{\text{minimize}} && \frac{1}{2} \|X - \Theta\|_F^2 \\ & \text{subject to} && \text{rank}(\Theta) = k \end{aligned}$$

40 where  $\|\cdot\|_F$  denotes the Frobenius norm. The optimal  $\Theta$  is a  $k$ -dimensional linear subspace  
 41 that can be written as the product of the projected observations  $A \in \mathbb{R}^{n \times k}$  and the basis  
 42  $V \in \mathbb{R}^{k \times d}$ :

$$X \approx \Theta = AV.$$

43 This suggests that each observation  $x_i \in \text{rows}(X)$  can be well-approximated by a linear  
 44 combination of  $k$  basis vectors (the rows of  $V$ ):

$$x_i \approx \theta_i = a_i V$$

45 for  $i = 1, \dots, n$ .

46 **Probabilistic Interpretation**

47 The PCA objective is equivalent to maximum likelihood estimation for a Gaussian model.  
 48 Under this lens, each observation  $x_i$  is a noisy realization of a  $d$ -dimensional Gaussian at  
 49  $\theta_i \in \text{rows}(\Theta)$ :

$$x_i \sim \mathcal{N}(\theta_i, I).$$

50 To recover the latent structure  $\Theta$ , PCA solves

$$\begin{aligned} & \underset{\Theta}{\text{maximize}} && \sum_{i=1}^n \log \mathcal{L}(x_i; \theta_i) \\ & \text{subject to} && \text{rank}(\Theta) = k \end{aligned}$$

51 where  $\mathcal{L}$  is the likelihood function.

52 **Exponential Family PCA**

53 **Link Function**

54 The link function  $g(\theta)$  connects the natural parameter  $\theta$  to the mean parameter  $\mu$  of an  
 55 exponential family distribution. It is defined as the gradient of the log-partition function  $G(\theta)$ :

$$\mu = g(\theta) = \nabla G(\theta).$$

56 The link function serves a role analogous to that in generalized linear models (GLMs) (McCullagh  
 57 & Nelder, 1989). In GLMs, the link function connects the linear predictor to the mean of  
 58 the distribution, enabling flexibility in modeling various data types. Similarly, in EPCA, the  
 59 link function maps the low-dimensional latent variables to the expectation parameters of  
 60 the exponential family, thereby generalizing the linear assumptions of traditional PCA to  
 61 accommodate diverse distributions (see [appendix](#)).

62 **Bregman Divergences**

63 EPCA extends the probabilistic interpretation of PCA using a measure of statistical difference  
 64 called the Bregman divergence (Bregman, 1967; Efron, 2004). The Bregman divergence  $B_F$   
 65 for a strictly convex, continuously differentiable function  $F$  is

$$B_F(p\|q) = F(p) - F(q) - \langle \nabla F(q), p - q \rangle.$$

66 This can be interpreted as the difference between  $F(p)$  and its linear approximation about  
 67  $q$ . When  $F$  is the convex conjugate of the log-partition function of an exponential family  
 68 distribution, minimizing the Bregman divergence corresponds to maximizing the associated  
 69 log-likelihood (Azoury & Warmuth, 2001; Forster & Warmuth, 2002) (see [documentation](#)).

70 **Loss Function**

71 EPCA generalizes the PCA objective as a Bregman divergence between the data  $X$  and the  
 72 expectation parameters  $g(\Theta)$ :

$$\begin{aligned} & \underset{\Theta}{\text{minimize}} && B_F(X\|g(\Theta)) \\ & \text{subject to} && \text{rank}(\Theta) = k \end{aligned}$$

73 where

- 74 ■  $g(\theta)$  is the **link function** and the gradient of  $G$ ,
- 75 ■  $G(\theta)$  is a strictly convex, continuously differentiable function (usually the **log-partition**  
 76 of an exponential family distribution),
- 77 ■ and  $F(\mu)$  is the **convex conjugate** of  $G$  defined by

$$F(\mu) = \langle \mu, \theta \rangle - G(\theta).$$

78 This suggests that data from the exponential family is well-approximated by expectation  
 79 parameters

$$x_i \approx g(\theta_i) = g(a_i V).$$

80 **Regularization**

81 To ensure the optimum converges, we introduce a regularization term

$$\begin{aligned} & \underset{\Theta}{\text{minimize}} && B_F(X\|g(\Theta)) + \epsilon B_F(\mu_0\|g(\Theta)) \\ & \text{subject to} && \text{rank}(\Theta) = k \end{aligned}$$

82 where  $\epsilon > 0$  and  $\mu_0 \in \text{range}(g)$ .

83 **Example: Poisson EPCA**

84 The Poisson EPCA objective is the generalized Kullback-Leibler (KL) divergence (see [appendix](#)),  
 85 making Poisson EPCA ideal for compressing discrete distribution data.

86 This is useful in applications like belief compression in reinforcement learning (Roy et al.,  
 87 2005), where high-dimensional belief states can be effectively reduced with minimal information  
 88 loss. Below we recreate a figure from Roy & Gordon (2002) and observe that Poisson EPCA  
 89 achieved a nearly perfect reconstruction of a 41-dimensional belief profile using just 5 basis  
 90 components.

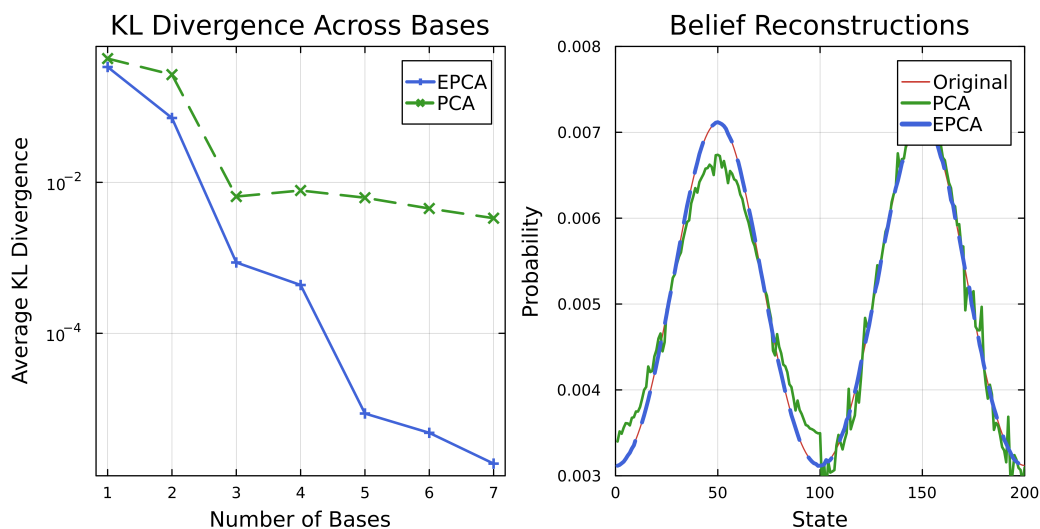


Figure 1: Left - KL Divergence for Poisson EPCA versus PCA. Right - Reconstructions from the models.

91 For a larger environment with 200 states, PCA struggles even with 10 basis.

## 92 API

### 93 Supported Distributions

94 ExpFamilyPCA.jl includes efficient EPCA implementations for several exponential family  
95 distributions.

Julia	Description
BernoulliEPCA	For binary data
BinomialEPCA	For count data with a fixed number of trials
ContinuousBernoulliEPCA	For modeling probabilities between 0 and 1
GammaEPCA	For positive continuous data
GaussianEPCA	Standard PCA for real-valued data
NegativeBinomialEPCA	For over-dispersed count data
ParetoEPCA	For modeling heavy-tailed distributions
PoissonEPCA	For count and discrete distribution data
WeibullEPCA	For modeling life data and survival analysis

### 96 Custom Distributions

97 When working with custom distributions, certain specifications are often more convenient and  
98 computationally efficient than others. For example, inducing the gamma EPCA objective from  
99 the log-partition  $G(\theta) = -\log(-\theta)$  and its derivative  $g(\theta) = -1/\theta$  is much simpler than  
100 implementing the full the Itakura-Saito distance (Itakura & Saito, 1968) (see appendix):

$$D(P(\omega), \hat{P}(\omega)) = \frac{1}{2\pi} \int_{-\pi}^{\pi} \left[ \frac{P(\omega)}{\hat{P}(\omega)} - \log \frac{P(\omega)}{\hat{P}(\omega)} - 1 \right] d\omega.$$

101 In ExpFamilyPCA.jl, we would write:

```
G(θ) = -log(-θ)
g(θ) = -1 / θ
gamma_epca = EPCA(indim, outdim, G, g, Val(:(G, :g))); options = NegativeDomain()
```

102 A lengthier discussion of the EPCA constructors and math is provided in the [documentation](#).

### 103 Usage

104 Each EPCA object supports a three-method interface: `fit!`, `compress`, and `decompress`. `fit!`  
105 trains the model and returns the compressed training data; `compress` returns compressed input;  
106 and `decompress` reconstructs the original data from the compressed representation.

```
X = sample_from_gamma(n1, indim)
```

```
Y = sample_from_gamma(n2, indim)
```

```
X_compressed = fit!(gamma_epca, X)
```

```
Y_compressed = compress(gamma_epca, Y)
```

```
Y_reconstructed = decompress(gamma_epca, Y_compressed)
```

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